GRADUATE STUDENTS' EFFECT SIZE CATEGORY BOUNDARIES

<u>V. N. Vimal Rao</u>¹, Jeffrey K. Bye¹, and Sashank Varma² ¹Department of Educational Psychology, University of Minnesota ²School of Interactive Computing & School of Psychology, Georgia Institute of Technology rao00013@umn.edu

Statisticians increasingly decry ritualistic categorizations of statistical measures. The interpretation of effect sizes is often guided by benchmarks, i.e., Cohen's d = .2 represents a 'small' effect size; .5 represents a 'medium' effect size; and .8 ('large') represents a large effect size. We employed a cognitive science approach to investigate how researchers systematically categorize values between these benchmarks. We find effect size categories are typically separated by fuzzy boundaries, as predicted by psychological theories of categorization. Understanding the cognitive processes underlying statistical reasoning can help us consider how to move beyond ritualistic interpretation of statistical measures.

INTRODUCTION

In 2019, Wasserstein, Schirm, and Lazar, on behalf of the American Statistical Association, called for an end to the era of statistical significance. As many fields have moved to emphasize effect sizes (e.g., Cumming, 2014), Wasserstein et al. additionally give a warning for the future: "to preclude a reappearance of this problem elsewhere, we must not begin arbitrarily categorizing other statistical measures" (Wasserstein et al., 2019, p. 2).

However, to cognitive scientists, categorization is fundamental to cognition—perception of a stimulus and seeing it *as something* is, at its heart, an act of categorization (Goldstone et al., 2013). Categorization naturally emerges whenever we respond differently to objects based on some attribute (Harnad, 1987), such as interpreting an effect size from its numerical magnitude.

In this paper we first consider why individuals might inherently categorize statistical measures. We then examine whether the widespread use of benchmarks underlies categorical interpretations of effect sizes in a manner antithetical to Wasserstein et al.'s (2019) warning against categorization.

BACKGROUND

Categorization is ubiquitous to cognition (Harnad, 1987). Thus, how might it present itself during acts of statistical thinking? Concepts are mental representations of categories; they denote what objects are being represented and how that information can be used to make inferences (Smith, 1989). Concepts provide structure to our interactions with the external world.

- Concepts efficiently encode information, reducing cognitive processing (e.g., Bruner et al., 1956; Goldstone et al., 2013). For example, rather than storing complete information about every right skew distribution one has encountered, one might only store a single representation of a prototypical *right skew* distribution.
- Concepts facilitate the generalization of experiences to objects within the same category (e.g., Goodman et al., 2008). For example, the concept of *multicollinearity* provides information regarding the interaction between two explanatory variables categorized as multicollinear in a statistical model.
- Individuals who share concepts can succinctly communicate information with one another (e.g., Markman & Makin, 1998). For example, describing a variable as a *confounder* (to someone who shares the concept of a confounder) communicates information about its relationship with other variables.

Cognitive scientists generally accept the ubiquity of concepts and view a wide variety of cognitive acts as fundamentally an act of categorization (Murphy, 2002).

Benchmarks and Boundaries

Benchmarks may serve as a representative 'ideal' for a concept. Benchmarks can either be explicitly specified (e.g., Cohen's d = .5 is a 'medium' effect size) or formed implicitly, for example, as a weighted average of all members of the category (i.e., prototype theory; Rosch, 1975). Individuals

use such benchmarks to determine category membership of novel stimuli based on the similarly of stimuli to the benchmark (Rosch & Mervis, 1975), often yielding a gradient effect and a 'fuzzy' boundary for what does or does not belong in the category.

For example, when determining whether a given empirical distribution is normal, statisticians may compare it to the mathematically defined normal distribution, which serves as the *benchmark* normal distribution. If a given distribution is similar enough to this benchmark, we may treat the distribution in the same way we would treat the benchmark normal distribution. As distributions start to deviate from this prototype (e.g., due to skew, kurtosis, or modality), decisions may become 'fuzzier.'

Even in the absence of pre-specified benchmarks, individuals implicitly build a notion of category *typicality* (represented by benchmarks) through repeated exposure to stimuli and the act of choosing how to respond to the stimuli (Posner & Keele, 1968). For example, when examining the normality of residuals from a simple linear regression model, a statistician decides whether the distribution is 'acceptable' or some remedial action must be taken. This helps them to form a cohesive concept of *normal* with which to categorize residual plots. Through such repeated exposure to stimuli and their associated responses, individuals become able to efficiently delineate distinct categories of stimuli by reference to benchmarks. Categorization researchers measure the degree to which a new stimulus is *similar* to a benchmark based on the extent to which the behavioral responses are similar (Palmeri, 1997).

In contrast to the fuzzy boundaries for concepts based on benchmarks, some concepts may be defined by the *boundary* around the category (Ashby, 1992). In these cases, the act of specifying a boundary leads to the formation of a concept. For example, 'p < .05' is a common boundary delineating 'statistical significance,' and according to the strict logic of null hypothesis significance testing, p = .051 is not meaningfully different from p = .273. When stimuli near a categorical boundary are treated similarly to those farther away, we may say the concept has a 'hard' boundary. In some cases, concepts defined by benchmarks can also result in a hard boundary (e.g., Davis & Love, 2010).

A New Statistics with Old Problems?

Much like the *p*-value controversy where 'statistical significance' created a publication bias against studies with large *p*-values, there is already evidence of a burgeoning effect size controversy replete with its own publication bias. For example, Schäfer and Schwarz (2019) found a problematic difference in the distribution of effect sizes between publications with pre-registration and those without. Is this because individuals are already categorizing effect sizes like they categorized *p*-values? Consistent with this possibility, Collins and Watt (2021) found that the overwhelming majority of psychology researchers they surveyed consider the values provided by Cohen (1988) as best exemplifying 'small,' 'medium,' and 'large' effect sizes, despite Cohen's warning that these values were arbitrarily chosen.

The reification of categories and concepts in the interpretation of statistical measures can alter the manner in which individuals perceive the measures. This can sometimes lead to a categorical perception effect, where perceived differences across categories are exaggerated and perceived differences within a category are diminished. These effects have been documented in the initial processing of *p*-values (Rao et al., 2022). Cohen's *d* effect sizes are typically defined by benchmarks: d = .2 ('small'), d = .5 ('medium'), and d = .8 ('large'). It is possible that through repeated instruction and practice, the widespread familiarity with Cohen's *d* benchmarks (Collins & Watt, 2021) may reinforce the cognitive concept of 'small,' 'medium,' and 'large' effect sizes, and this may in turn induce fuzzy or hard boundaries between effect size categories.

It is currently unknown how researchers systematically categorize Cohen's *d* values falling between these effect size benchmarks. Therefore, the purpose of this study is to examine, through implicit and explicit measures, where and how researchers draw boundaries between effect size categories: at what magnitude does an effect size 'change' from being categorized as 'small' to 'medium' and from 'medium' to 'large,' and is this change gradual (as with a fuzzy boundary) or immediate (as with a hard boundary)?

METHODS

To identify the location of boundaries between effect size categories, we employed a cognitive science approach. Boundary identification tasks are commonly used as a first step in evaluating the

cognitive effects of categories and concepts on individuals' interactions with stimuli, especially the *implicit* boundaries between categories (e.g., at which hue(s) do individuals switch from labeling hues as 'blue' to labeling them as 'green'?).

Graduate students in the psychological sciences at a research university in the Midwestern United States were recruited for this study (n = 39). All participants had completed at least one year of instruction and training in statistical methods at the doctoral level. They completed the boundary identification task as the second of three tasks. The full study took approximately 40 minutes on average to complete in full, and participants were compensated with a \$25 electronic gift card.

Participants were first reminded that Cohen's d is "a statistic indicating the size of an effect in standard deviation units." They were then told that they would be shown various values of Cohen's d and for each value, they would judge whether it indicated 'no effect,' a 'small effect,' a 'medium effect,' or a 'large effect,' by selecting one of four keys on a keyboard. Crucially, participants were not told how to make this judgment, and at no point in the study were the standard benchmarks (i.e., .2, .5, and .8) mentioned to participants.

Participants completed 180 trials in four blocks of 45 stimuli each, preceded by eight practice trials. There were 90 unique stimuli of the form "d = .XX" with values ranging from .01 to .90. Participants categorized each value twice: once in the first two blocks and again in the last two. Within each set of two blocks, the stimuli order was randomly shuffled. Stimuli were presented one-at-a-time and remained on screen until participants made their selection. Participants were encouraged to make their initial selection as quickly as possible.

After completing the full study, participants also completed a short survey collecting basic demographic information and probing for possible demand characteristics for the study. As part of this survey, participants were asked to *explicitly* specify the upper and lower bounds of what they would consider a 'small,' 'medium,' and 'large' Cohen's *d* effect size.

RESULTS

To identify the point at which participants' effect size categories 'changed' from being categorized as 'small' to 'medium' and from 'medium' to 'large,' we first analyzed their survey responses. Of the 39 participants, 34 explicitly self-reported that they referred to the values of .2, .5, and .8 as *benchmarks* for categorizing effect sizes. The remaining five participants referred to the values .1, .3, and .5—common benchmarks in the interpretation of correlation coefficients (i.e., r). Data from these five participants were analyzed separately.

Participants varied in their explicit *boundaries* between the categories, with the median boundary delineating 'small' and 'medium' effect sizes at .39, and the median boundary delineating 'medium' and 'large' effect sizes at .70. Surprisingly, very few participants specified categorical boundaries at the arithmetic midpoint (i.e., mean) between common benchmark values (i.e., .35 and .65; see the left panel of Figure 1). This may reflect that they had a variety of interpretations of the benchmark values of .2, .5, and .8. Participants drawing a boundary between 'small' and 'medium' effect sizes near .5 might have interpreted .5 as a boundary rather than a benchmark, as is typical of other statistical measures such as *p*-values (where .05 serves as a categorical boundary). Those drawing the same boundary near .3 might be influenced by the desirability of finding a 'medium' effect rather than a 'small' one. Those drawing the boundary near .4 might adhere to a conservative approach based on an aversion to taking the risk of over-interpreting statistical results and possibly committing a questionable research practice.

At the aggregate level, the implicit boundaries from participants' responses on the boundary identification task matched their explicit boundaries from the survey. As seen in the right panel of Figure 1, participants' responses showed average implicit categorical boundaries at .38 (delineating 'small' and 'medium' effect sizes) and .69 (delineating 'medium' and 'large' effect sizes), consistent with the median explicit values. Interestingly, there is quite a bit of overlap in the assigned labels for a given effect size. This may be due to the psychological boundary between effect size categories indeed being fuzzy, or due to variability in the location of hard boundaries amongst participants, as observed in the explicit boundary values.



Figure 1. Aggregate response patterns: explicit boundary location survey responses with reference lines at aggregate boundaries (left) and boundary identification task responses with reference lines at aggregate implicit boundaries (right)

Participants categorized each effect size between .01 and .90 twice, but their two responses were not always in agreement (see the left panel of Figure 2). The average agreement rate across all participants and effect size values was approximately 88%. The agreement rate was lower (as low as 60%) near implicit boundaries (i.e., values of .11, .38, and .69, as identified in the aggregate response selections), and higher near benchmarks (i.e., values of .2, .5, and .8). Interestingly, the locations of the *most* consistently categorized effect sizes were not the benchmark values themselves, but rather at slightly higher values.



Figure 2. Response patterns: consistency in response selections (left) and average response time (right) by effect size value with reference lines at aggregate boundaries

A complementary pattern is also seen in participants' response times (see the right panel of Figure 2). Participants' response times in selecting a category label were approximately 12% slower when categorizing values near implicit boundaries relative to benchmark values. Participants were fastest in making their selection at values slightly higher than the benchmark values. These response patterns exhibit typicality effects consistent with a 'fuzzy' boundary (Rosch & Mervis, 1975).

An examination of individual participants' response profiles revealed that most participants' (23 of 39) response patterns clearly reflected a fuzzy boundary between categories, as exemplified by the response profile of GID22 (see Figure 3, bottom right). These profiles exhibit overlap between categories as well as increased response times and decreased consistency near category boundaries.

Some participants demonstrated hard boundaries, especially the 11 participants who interpreted common benchmark values as boundaries, rather than drawing boundaries between benchmark values. For example, the response profile of participant GID1 (Figure 3, top left) indicates a hard boundary between 'medium' and 'large' Cohen's *d* effect sizes at .5. Similarly, .5 serves as a hard boundary between 'small' and 'medium' Cohen's *d* effect sizes for participant GID38 (Figure 3, top right).



Figure 3. Selected participant response patterns: participants' responses and response time (RT) with reference lines at common benchmark values as provided by Cohen (1988)

However, not all categorical boundaries were hard boundaries for these participants. The boundary between 'small' and 'medium' effect sizes appears to be a fuzzy boundary for GID1, evidenced by the overlap in the category labels assigned to each effect size. Similarly, the boundary between 'medium' and 'large' effect sizes is a fuzzy boundary for GID38.

Some participants who interpreted the values of .2, .5, and .8 as benchmarks still drew hard boundaries between categories. For example, participant GID26's response profile (Figure 3, bottom left) shows the benchmarks relatively centered within each category, and hard boundaries between categories at .11, .40, and .68 respectively. However, GID26's response times were generally longer for effect size values near their category boundaries than for values near category benchmarks, consistent with a 'fuzzier' boundary.

DISCUSSION

In this study we investigated how researchers categorize effect sizes into the commonly utilized 'small,' 'medium,' and 'large' categories using both implicit and explicit measures. We found that effect size categories were typically (but not always) separated by 'fuzzy' boundaries, as predicted by psychological theories of categorization.

Surprisingly, participants' implicit and explicit response patterns indicate they do not draw boundaries *exactly* at the arithmetic midpoints between common benchmark values, nor are they fastest and most consistent at *exactly* the common benchmark values. This may be due to the way in which we perceive symbolic (and non-symbolic) numbers. The standard model of numerical cognition suggests we possess a logarithmically compressed mental number line with psychological boundaries based on our place value system (Moyer & Landauer, 1967; Nuerk et al., 2011; Varma & Karl, 2013). Therefore, participants' boundaries may reflect psychological midpoints based on their mental number line.

This study is the first to empirically explore how researchers categorize a wide range of effect sizes, specifically in how they draw boundaries between effect size categories. Using labels such as those commonly utilized for Cohen's d effect sizes affects not only students' benchmarks but also the boundaries between them, sometimes in unpredicted ways. Understanding the cognitive processes underlying statistical reasoning can inform what we should practice and what we should teach if we are to move beyond the ritualistic categorical interpretation of statistical measures.

REFERENCES

- Ashby, F. G. (1992). *Multidimensional models of perception and cognition*. Erlbaum.
- Bruner, J. S., Goodnow, J. J., & Austin, G. A. (1956). A study of thinking. Wiley. https://doi.org/10.2307/1292061
- Cohen, J. (1988). Statistical power analysis for the behavioral sciences (2nd ed.). Erlbaum.
- Collins, E., & Watt, R. (2021, July 1). Use, knowledge, and misconceptions of effect sizes in psychology. PsyArXiv. <u>https://doi.org/10.31234/osf.io/r7vmf</u>
- Cumming, G. (2014). The new statistics: Why and how. *Psychological Science*, 25(1), 7–29. <u>https://doi.org/10.1177/0956797613504966</u>
- Davis, T., & Love, B. C. (2010). Memory for category information is idealized through contrast with competing options. *Psychological Science*, 21(2), 234–242. https://doi.org/10.1177/0956797609357712
- Goldstone, R. L., Kersten, A., & Carvalho, P. F. (2013). Concepts and categorization. In A. F. Healy, R. W. Proctor, & I. B. Weiner (Eds.), *Handbook of psychology: Experimental psychology* (pp. 607–630). John Wiley & Sons. <u>https://doi.org/10.1002/9781119170174.epcn308</u>
- Goodman, N. D., Tenenbaum, J. B., Feldman, J., & Griffiths, T. L. (2008). A rational analysis of rulebased concept learning. *Cognitive Science*, 32(1), 108–154. https://doi.org/10.1080/03640210701802071
- Harnad, S. (1987). Psychophysical and cognitive aspects of categorical perception: A critical overview. In S. Harnad (Ed.), *Categorical perception: The groundwork of cognition* (pp. 1–25). Cambridge University Press.
- Markman, A. B., & Makin, V. S. (1998). Referential communication and category acquisition. *Journal* of Experimental Psychology: General, 127(4), 331–354. <u>https://doi.org/10.1037/0096-3445.127.4.331</u>
- Moyer, R. S., & Landauer, T. K. (1967). Time required for judgments of numerical inequality. *Nature*, 215(5109), 1519–1520. <u>https://doi.org/10.1038/2151519a0</u>
- Murphy, G. L. (2002). The big book of concepts. MIT Press.
- Nuerk, H. C., Moeller, K., Klein, E., Willmes, K., & Fischer, M. H. (2011). Extending the mental number line. *Zeitschrift für Psychologie*, 219(1), 3–22. <u>https://doi.org/10.1027/2151-2604/a000041</u>
- Palmeri, T. J. (1997). Exemplar similarity and the development of automaticity. Journal of Experimental Psychology: Learning, Memory, and Cognition, 23(2), 324–354. <u>https://doi.org/10.1037/0278-7393.23.2.324</u>
- Posner, M. I., & Keele, S. W. (1968). On the genesis of abstract ideas. Journal of Experimental Psychology, 77(3, Pt.1), 353–363. <u>https://doi.org/10.1037/h0025953</u>
- Rao, V. N. V., Bye, J. K., & Varma, S. (2022). Categorical perception of *p*-values. *Topics in Cognitive Science*, 14(2), 414–425. <u>https://doi.org/10.1111/tops.12589</u>
- Rosch, E. (1975). Cognitive representations of semantic categories. *Journal of Experimental Psychology: General, 104*(3), 192–233. <u>https://doi.org/10.1037/0096-3445.104.3.192</u>
- Rosch, E., & Mervis, C. B. (1975). Family resemblances: Studies in the internal structure of categories. *Cognitive Psychology*, 7(4), 573–605. <u>https://doi.org/10.1016/0010-0285(75)90024-9</u>
- Schäfer, T., & Schwarz, M. A. (2019). The meaningfulness of effect sizes in psychological research: Differences between sub-disciplines and the impact of potential biases. *Frontiers in Psychology*, 10, Article 813. <u>https://doi.org/10.3389/fpsyg.2019.00813</u>
- Smith, E. E. (1989). Concepts and induction. In M. I. Posner (Ed.), *Foundations of cognitive science* (pp. 501–526). MIT Press.
- Varma, S., & Karl, S. R. (2013). Understanding decimal proportions: Discrete representations, parallel access, and privileged processing of zero. *Cognitive Psychology*, 66(3), 283–301. <u>https://doi.org/10.1016/j.cogpsych.2013.01.002</u>
- Wasserstein, R. L., Schirm, A. L., & Lazar, N. A. (2019). Moving to a world beyond "*p* < 0.05." *The American Statistician*, 73(S1), 1–19. <u>https://doi.org/10.1080/00031305.2019.1583913</u>